The impact of group membership on punishment versus partner rejection

Trystan Loustau\*, Jacob Glassman, Justin W. Martin, Liane Young, Katherine McAuliffe

Department of Psychology and Neuroscience, Boston College, Chestnut Hill, MA

Correspondence: Trystan Loustau, loustau@bc.edu

Address: 275 Beacon St, Chestnut Hill, MA 02467

**Abstract**

People often display ingroup bias in punishment, punishing outgroup members more harshly than ingroup members. However, the impact of group membership may be less pronounced when people are choosing whether to stop interacting with someone (i.e., partner rejection). In two studies (N = 1667), we investigate the impact of group membership on both response types. Participants were assigned to groups based on a “minimal” groups paradigm (Study 1) or their self-reported political positions (Study 2) and played an incentivized economic game with other players. In this game, participants (Responders) responded to other players (Deciders). In the Punishment condition, participants could decrease the Decider’s bonus pay. In the Partner Rejection condition, participants could reject future interactions with the Decider. Participants played once with an ingroup member and once with an outgroup member. To control for the effects of intent and outcome, scenarios also differed based on the Decider’s Intent (selfish versus fair) and the Outcome (equal versus unequal distribution of resources). Participants punished outgroup members more than ingroup members, however group membership did not influence decisions to reject partners. These results highlight partner rejection as a boundary condition for the impact of group on responses to transgressions.**Introduction**

Group membership influences a variety of behaviors, including cooperation, conflict, empathy, trust, and many others [1][2][3][4]. Of recent interest is the tendency for group membership to influence how people respond when others transgress against them (e.g., when they harm them or others, violate a norm, or otherwise behave antisocially). Much of this work has focused on one type of response: punishment—the imposition of a cost on a transgressor [5][6][7][8]. Past research has found that outgroup members are often punished more harshly than ingroup members by third-party observers [9][10][11][12][13]. However, in other cases, third-party observers may instead punish ingroup members more harshly than outgroup members [14]. While work on third-party punishment points to an important role for group membership, these data do not speak to how individuals react when they themselves are the victims of the transgression. This distinction is crucial because the dynamics of personal involvement in an offense may lead to different patterns of punitive behavior compared to third-party punishment.

Relatively less work has examined ingroup bias in an interpersonal context (i.e., second-party punishment). However, what work exists has yielded mixed findings, with some studies showing that people punish outgroup members more harshly [15], other studies showing that people punish ingroup members more harshly [16], and yet other studies showing no effect of group membership on punishment in an interpersonal context [17]. To help clarify these effects, we examine the impact of group membership on second-party punishment decisions in the present work.

Importantly, punishment is not the only way people can respond to others’ transgressions. They may also respond with partner rejection, refusing to interact with the transgressor any further [18][19]. Similar to punishment, partner rejection is an important and frequently employed method for navigating social relationships and social conflict [20]. Partner rejection can be observed in many areas of life, such as in the termination or rejection of friendships, romantic relationships, employment opportunities, roommate and housing situations, or club memberships. Thus, it is important to consider the extent to which partner rejection is influenced by group membership. Examining the sensitivity of partner rejection to group membership is also critical to clarifying the boundary conditions for the impact of group membership status on responses to transgressions. In the present work, we investigate the influence of group membership on both punishment and partner rejection in response to interpersonal transgressions.

**Distinguishing Punishment and Partner Rejection**

Traditional accounts of interpersonal punishment, that is, punishment inflicted upon a transgressor by the victim of a transgression, have defined it as the imposition of a cost on a transgressor to change a behavior [21]. These costs can vary; they may be monetary (e.g., pay cut or fine), physical (e.g., corporal punishment), or psychological and emotional (e.g., reputational damage). Modern accounts of punishment suggest that punishment serves a communicative function, demonstrating that people infer that punishment will be effective if it communicates social and moral norms even when it poses no cost to the transgressor [22]. This account is supported by prior work showing that punishment signals and enforces moral values and social rules [23][24][5][25][26][27]. Given the communicative role of punishment, it makes sense that punishment is sensitive to group membership.

Punishing outgroup members more harshly than ingroup members communicates lower tolerance for mistreatment by outgroup members than ingroup members [28]. As demonstrated by prior work, the impact of group membership on punishment is flexible and responsive to the communicative goals of the punisher. For instance, ingroup members are often punished more harshly than outgroup members in the context of group deviance, communicating lower tolerance for such deviance by ingroup members than outgroup members [16].

Most research relevant to the function of partner rejection focuses on *partner choice* broadly, which includes both partner *selection* (i.e., choosing to interact with someone) and partner *rejection* (i.e., choosing not to interact with someone or not to continue interacting with someone). Indeed, some work using the term *partner choice* measures only one of the two components—partner selection or partner rejection [29]. For clarity, we discuss relevant literature on partner choice, but primarily use the term partner rejection, as that was the focus of the present work.

Within the traditional accounts of punishment outlined above, partner rejection might be classified as a form of punishment, one that imposes costs through the severance of a social tie (e.g., emotional pain, loss of resources, and social connection). Yet, in contrast to punishment, partner choice is theorized to serve as a tool for navigating the biological market [30]. Partner choice decisions are sensitive to market dynamics, as partner selectors and rejectors aim to maximize their personal benefits, creating pressure for potential partners to make themselves look appealing to partner selectors and for current partners to behave cooperatively [18][19]. As evidence of this, participants tend to behave more cooperatively when participants can choose or reject their partners in economic games [31][32].

Given the market function of partner choice decisions for maximizing one’s own benefits, partner choice may only be sensitive to group membership to the extent that group membership provides information regarding the impact that a particular partnership (new or continued) will have on personal well-being. For example, people may be more likely to choose interaction partners from ingroups that are trusted to be more cooperative than outgroups that are not. However, when group membership does not provide much information about a potential partner’s generosity and fairness, it may not be factored into partner choice decisions. Indeed, one study found that participants’ choices for a partner before an economic game were not sensitive to whether they belong to the same ethnic or gender group but were sensitive to how reliable they perceived them to be based on a photograph [33]. Similarly, people’s decisions to reject a partner who has mistreated them may be insensitive to the partner’s group membership, given that group membership is unrelated to the mistreatment.

This account of partner choice, that emphasizes individual well-being, is consistent with prior work showing that second-party partner choice decisions are more strongly influenced by the intentions of interaction partners (i.e., to be fair or selfish) than the outcome (e.g., unequal or equal distribution of resources) of the interactions [34][35]. Seeking to maximize their individual well-being, people may reject partners who intended to be selfish even if the outcome was equal in order to protect themselves from future mistreatment. By contrast, the same work shows that decisions to punish an interaction partner are relatively equally influenced by their intentions and the outcome. This balanced consideration of intentions and outcomes supports the role of punishment in both addressing the immediate injustice and signaling norms and expectations for future interactions.

Both partner rejection and punishment represent ways of responding to an individual transgressor. In the present work, we investigate the impact of group membership on partner rejection and punishment in response to individual transgressions when group membership is not inherently tied to the transgressions. We examine the impact of group membership on the likelihood of partner rejection and the likelihood of punishment separately, as well as directly compare the strength of impact of group membership on these two types of responses. Building on previous work highlighting that punishment and partner rejection are differentially sensitive to intent and outcome, we designed studies that allow us to incorporate those differences while investigating the impact of group membership.

**Current Studies**

In the present work, we investigate the impact of group membership on second-party decisions to punish or reject a partner. Across two studies, we examine participants’ punishment and partner rejection decisions in an online economic game adopted from past work. We manipulate group membership using a multi-pronged approach, through a “minimal” groups paradigm (Study 1) and based on a self-reported political opinion (Study 2). In addition, in order to account for differences in punishment and partner rejection that may be due to the perceived intent of the partner and the outcome of the interaction, intent and outcome varied within-subjects.

**Study 1**

## Method

### Participants

American participants (n = 802) were recruited via Amazon Mechanical Turk and completed an online survey in exchange for a small payment (≤ $2.00) with an opportunity to earn a bonus payment. We preregistered a target minimum sample size of n = 300 per condition. We excluded data from participants who failed comprehension questions or attention checks, who were not native English speakers, or whose mean reaction time was less than three standard deviations below the overall mean. These criteria yielded a final sample of 659 participants (17.8% excluded; final sample 48% male, 51% female, <0.5% other, <0.5% preferred not to answer). The percentage excluded was similar across both response type conditions (Punishment: 14.9%; Partner Rejection: 11.3%).

Post-hoc sensitivity analyses were conducted using the simr package [36] in R Studio to determine the minimum effect sizes that we were powered to detect in a linear mixed effects model and a logistic regression with the current sample. Using simr sensitivity analyses with 1000 simulations, we found that the linear mixed effects model was 81.00% (95% CI = 71.93, 88.16) powered to detect a three-way interaction of GroupXIntentXOutcome with an unstandardized beta of 0.13, and that the logistic regression model was 80.60% (95% CI = 78.01, 83.01) powered to detect a three-way interaction of GroupXIntentXOutcome with an unstandardized beta of 1.74 (OR = 5.70).

***Procedures***

First, participants were assigned to one of two minimal groups. Then, they interacted with other participants in an asynchronous online economic game modeled on past work [37][35]. Participants were randomly assigned to play one of two versions of the game based on response type: Punishment or Partner Rejection.

**Group assignment.**In Study 1, we employed a variant of the “minimal” groups paradigm **[2][38]** to assign group membership. After providing consent, participants were told they would perform a task that would determine which team they were on for the economic game. Participants were presented with a 3x3 word search and told to enter the first three-letter word they saw into a free response box. Participants were then told which team they belonged to based on the word they entered, TEAM CAT or TEAM OWL, and selected one of five potential cat or owl avatars to represent them throughout the task. To promote group identification, we informed participants that “people on the same team tend to share many personality and cognitive traits.” This paradigm assigned participants to groups about which they know almost nothing, allowing us to investigate the influence of the abstract group membership on punishment and partner rejection.

**Economic game.** Participants were then given instructions for the economic game. This game involved two roles: The Decider and the Responder. First, participants played as the Responder. On each trial, the Decider was required to choose one of two different methods for allocating $1 between themselves and the Responder: Option A and Option B. Option A resulted in a 2/3 chance of giving the full $1 to the Decider and a 1/3 chance of giving $0.50 to the Decider and $0.50 to the Responder. Option B resulted in a 2/3 chance of giving $0.50 to the Decider and $0.50 to the Responder and a 1/3 chance of giving the full $1 to the Decider. Thus, these options were designed such that Option A was more likely to yield an unequal outcome than Option B, but both options could have yielded an equal or unequal outcome. In this way, the Decider’s intent was dissociable from the outcome, resulting in four potential interactions: selfish intent and unequal outcome, selfish intent and equal outcome, fair intent and unequal outcome, and fair intent and equal outcome.

Participants played the game twice as a Responder: once with an ingroup member and once with an outgroup member in a randomized order. For each trial, participants were presented with all four possible combinations of the Decider’s choice and outcome (Table 1) and were asked to commit to a response for each option.

Participants were randomly assigned to one of two response conditions: Punishment or Partner Rejection. In the Punishment Condition, participants had the opportunity to add or subtract up to $0.30 in $0.10 increments from the Decider’s payoff. In both conditions, the decision to change Deciders or to punish was always costless. In the Partner Rejection Condition, participants were given a binary choice to either play another round with the same Decider or to play another round with a new Decider (i.e., reject the current decider). If the participant decided to reject their current Decider, the new Decider would always be from the same group as their current Decider (e.g., if they rejected a Decider from TEAM CAT, the Decider would be replaced with another Decider from TEAM CAT). In this way, neither the partner rejection nor punishment response could change the group which a participant’s Decider was from.

In both conditions, participants responded to all four potential outcomes. This method achieved two things: (1) we gained four times the amount of data than we would have using “hot” decisions, and (2) we could ask for participants’ decisions without informing them of the Decider’s decision or outcome on that round, so that the decisions between rounds were not influenced by the decisions or outcomes of previous rounds.

**Rating group identification.** After playing the game with both an ingroup and outgroup player, participants completed identification ratings for their ingroup and outgroup in a randomized order. We assessed these ratings after participants’ economic game participation to avoid the potential demand effects of assessing identification immediately following group assignment. Participants rated their agreement with four statements drawn from prior work (e.g., “I like TEAM [Owl/Cat]”; [39]) on a 100-pt slider scale from “Strongly disagree” to “Strongly agree”. These ratings had high reliability (Ingroup: Cronbach’s alpha = 0.92, 95% CI = 0.90–0.92; Outgroup: Cronbach’s alpha = 0.82, 95% CI = 0.80–0.84), so the ratings for each group were averaged for each participant.

**Additional Measures.** After playing as Responder,participants were told they would play this game two more times in the role of Decider. This provided us with Deciders to match with Responders. Procedures and findings for participants’ decisions as Deciders are reported in the Supplementary Materials. At the end of the study, participants completed eight attention check items, completed an optional demographics questionnaire, and were debriefed.

**Analysis approach.** All analyses were conducted in RStudio. Since responses in the Punishment Condition were on an interval scale (-$0.30 to $0.30 by $0.10) and responses in the Partner Rejection Condition were binary (reject or keep), it was not possible to compare these conditions in the same regression model. Therefore, we used linear mixed effects regression to analyze data from the Punishment Condition and logistic regression to analyze data from the Partner Rejection Condition. Group Membership (Ingroup versus Outgroup), Decider’s Intent (Fair versus Selfish), and Outcome (Equal versus Unequal) were entered as predictors in the full models. We also conducted exploratory analyses using a classification approach to determine the proportion of participants who were sensitive to intentions, outcomes, and group membership in each condition. Results for these exploratory analyses for each study are reported in the Supplementary Materials.

**Results and Discussion**

### Manipulation Check

Participants identified more strongly with their ingroup than their outgroup in both response type conditions, *t*(658) = 24.37, *p* < .001, *d* = 1.34. We found no difference between conditions in identification with ingroup members, *t*(654.63) = 1.41, *p* = .16, *d* = 0.11, or outgroup members, *t*(649.55) = 0.69, *p* = .49, *d* = 0.05. Thus, our group manipulation was successful in both response type conditions.

***Punishment Condition***

We conducted a series of linear-mixed effects regression with the lme4 package [40] to analyze data from the Punishment Condition. The models included a random intercept for each participant as well as random slopes for group membership, intent, and outcome. We used a model comparison approach by evaluating the importance of predictors using Likelihood Ratio Tests. First, we compared the full model with the three-way interaction term (Group MembershipXIntentXOutcome) to the same model without the three-way interaction term. The three-way interaction term did not improve model fit, (LRT *X*2 (1) = 0.01, *p* = .932). Next, we compared the model will all two-way interaction terms to the same model without the two-way interaction terms. We found that the two-way interaction terms improved model fit (LRT *X*2 (3) = 12.53, *p* = .006) (Table 2).

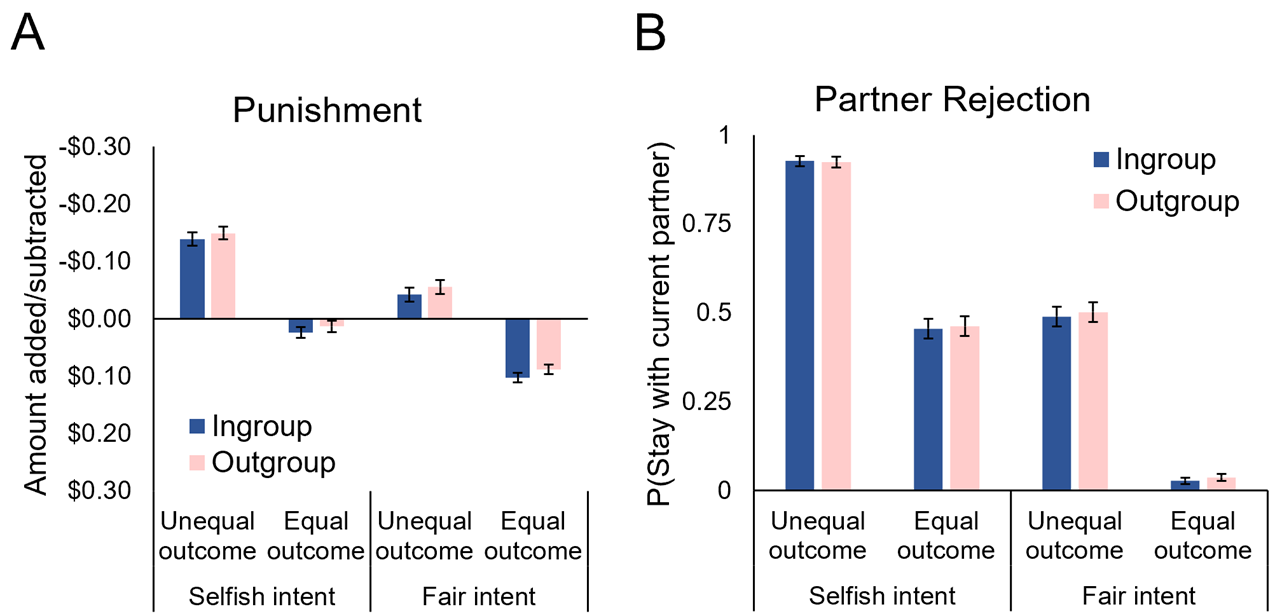
We found that there was a significant main effect of Group Membership (Table 3) such that participants punished less (i.e., subtracted less from the Decider’s payoff) when the Decider was an ingroup member than when they were an outgroup member. There was a significant main effect of Intent such that participants punished less when the Decider had a Fair intent than when they had a Selfish intent. Additionally, there was a significant main effect of Outcome such that participants punished less when the Outcome was Equal than when it was Unequal. There was a significant interaction between Intent and Outcome (Table 4) such that participants punished fair Deciders less than selfish Deciders, especially when the outcome was unfair than when the outcome was fair.

**Partner Rejection Condition.** First, we conducted a logistic mixed effects regression to analyze data in the Partner Rejection Condition, with a random intercept entered for each participant. However, results indicated that there was zero variance for this random intercept, so we reverted to standard logistic regression. Results showed that, in contrast to the Punishment Condition, there was no significant main effect of group membership (Table 5). There was a significant main effect of Intent such that participants were more likely to reject the Decider when they had a selfish intent than when they had a fair intent. There was also a significant main effect of Outcome such that participants were more likely to reject the Decider when there was an unequal outcome than when there was an equal outcome. Additionally, there was a significant interaction between Intent and Outcome such that participants were more likely to reject Deciders with a selfish intent than those with a fair intent, especially when the outcome was unequal. The effects of Intent and Outcome were consistent with the findings from the Punishment Condition.

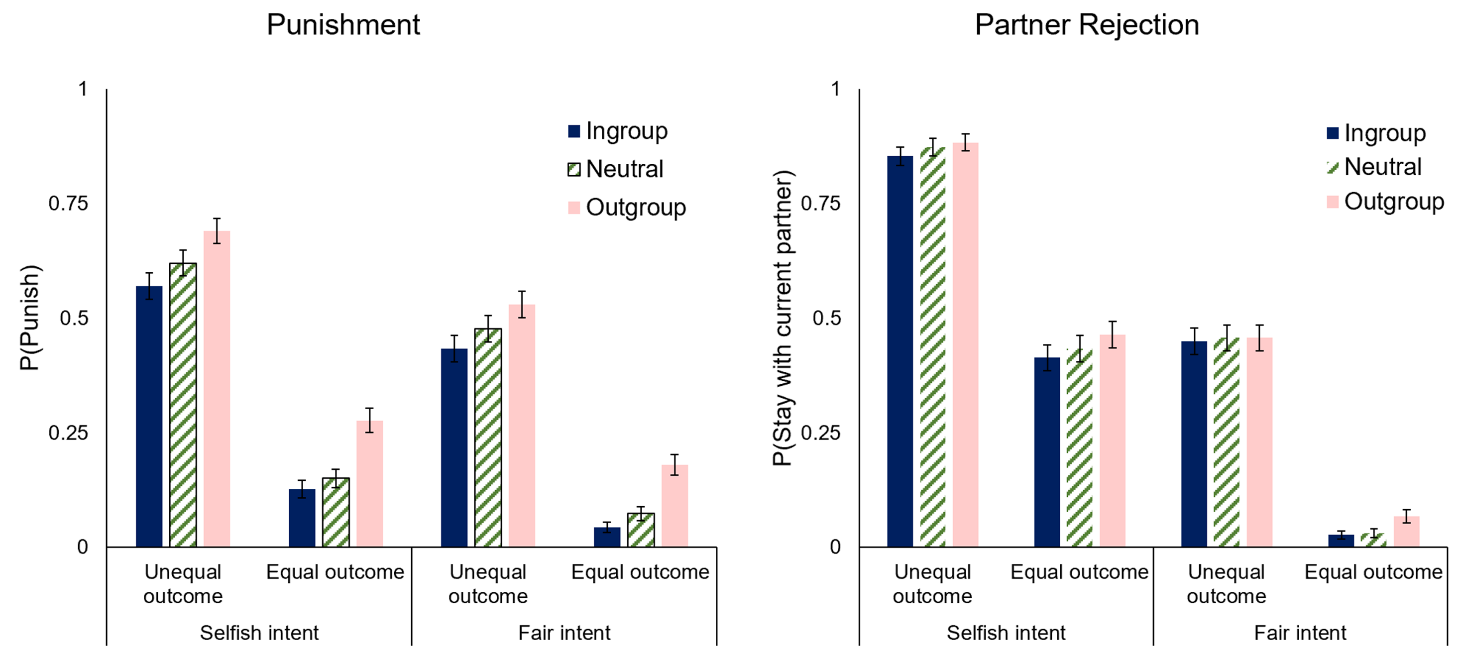
**Figure 1.**

Responses in the punishment and partner rejection conditions.

1. Study 1



1. Study 2



# Study 2

In Study 1, we found that participants punished outgroup members more harshly than ingroup members but did not reject outgroup members more than ingroup members. While the minimal groups paradigm we utilized in Study 1 allowed us to examine the impacts of abstract group membership, the lack of sensitivity to group boundaries in the Partner Rejection Condition may reflect the relatively weak nature of the minimum groups manipulation. To address this concern in Study 2, we assigned participants to groups based on their position on a consequential political issue (i.e., abortion). Assigning groups in this way invokes more salient boundaries which are frequently relevant in daily life. Utilizing these two approaches to group assignment allowed us to combine the strengths of both to obtain a better understanding of how group membership impacts responses to interpersonal transgressions, in line with prior work [41].

Additionally, in Study 2, we measured both punishment and partner rejection decisions on a binary scale, allowing us to analyze data from both conditions within the same regression model. We also included a third type of group member with which participants interacted: a neutral individual (who did not feel strongly about the political issue), which allowed us to examine whether the ingroup bias we observed in the Punishment Condition in Study 1 is more reflective of ingroup love (i.e., preferential treatment of an ingroup member compared to non-ingroup members) or outgroup derogation (i.e., preferential treatment of non-outgroup members compared to outgroup members).

In sum, Study 2 built upon the results of Study 1 by 1) examining the impact of group membership on punishment and partner rejection in a real-world group context, 2) directly comparing punishment and partner rejection by measuring them both on the same scale, and 3) comparing the influence of ingroup and outgroup membership to a control (neutral group membership).

## Method

### Participants

American participants (n = 865) were recruited via Amazon Mechanical Turk and completed an online survey in exchange for a small payment (≤ $2.00) with an opportunity to earn a bonus payment. We preregistered a target minimum sample size of n = 300 per condition. We excluded data from participants based on the same exclusion criteria used in Study 1 in addition to data from those who reported no position on the political issue used to assign participants to a group. These criteria yielded a final sample of 600 participants (23.2% excluded; final sample 47% male, 52% female, <0.5% other, <0.5% preferred not to answer). The percentage excluded was very similar across our response type conditions (Punishment: 22.7%; Partner Choice: 23.8%).

A post-hoc sensitivity analysis was conducted using the simr package [42] in R Studio to determine the minimum effect size (Odd’s Ratio) that this study was powered to detect in a two-way interaction, as our main analysis examined the interaction between group membership and response type (punishment versus partner choice). This study had .80 power to detect an effect size of at least 0.62 or 1.60 with a 5% error rate.

### Procedures

In Study 2, group membership was determined based on opinions about a political issue and participants played a modified version of the economic game used in Study 1.

**Group assignment.** After providing consent, participants were asked a series of standard demographic questions, and about their stance on abortion. Participants could identify as “Pro-Choice”, “Pro-Life” or “No position” and rated how strongly they supported this position on a Likert scale from 1 = “Not strongly at all” and 7 = “Very strongly”. Participants and the other players they interacted with were subsequently identified with an image corresponding to their choice.

**Economic game.** Participants played the same economic game from Study 1, with a few modifications. First, participants played the game three times instead of twice, once with a player from each group (Pro-life, Pro-choice, Neutral). The order of these interactions was counterbalanced across participants. Second, unlike the continuous measure of punishment in Study 1, there were only two response options available in the Punishment Condition. Since the most frequently chosen punishment options in Study 1 were to not punish (41% of responses) and maximum punishment (removing $0.30; 27% of responses), we used these as the punishment options available in Study 2.

**Rating group identification.** After playing the game with all three players, participants completed the same four-item measure of group identification used in Study 1 for each of the three groups. These ratings had high reliability (Ingroup: 0.91, 95% CI = 0.89–0.91; Outgroup: 0.92, 95% CI = 0.91–0.93; Neutral: 0.89, 95% CI = 0.88–0.90) and so the ratings for each group were averaged for each participant.

**Additional Measures.** As in Study 1, participants also played the game three times as the Decider. In an exploratory manner, the 7-item Cognitive Reflection Test (CRT; [43]) was administered after the game. Results for the Decider role and the CRT are reported in the Supplementary Materials. At the end of the study, participants completed eight attention check items and were debriefed.

**Analysis approach.** As specified in our pre-registration, we used mixed-effects logistic regression to examine the impact of response type and group membership on response. Selecting punishment or partner rejection was coded as 1, and selecting no punishment or partner rejection was coded as 0. Response Type (Punishment versus Partner Rejection), Group Membership (Ingroup versus Outgroup versus Neutral), Intent (Selfish versus Fair), and Outcome (Unequal versus Equal) were entered as fixed effects, as well as all possible interactions between these variables. We also included a random intercept for participants. As in Study 1, we used a model comparison approach to evaluate our data using Likelihood Ratio Tests.

**Results and Discussion**

### Manipulation Check: Did participants identify more with their ingroup?

We found that participants in both conditions identified with their ingroup significantly more than their outgroup or neutral individuals (Table 8). Identification with each group did not differ between response type conditions. Thus, our group manipulation successfully led participants to identify more with their own group compared to the two other groups.

### Primary Results

First, we compared the full model with the four-way interaction term (ResponseTypeXGroup MembershipXIntentXOutcome) to the same model without the four-way interaction term. The four-way interaction term did not improve model fit (LRT *X*2 (2) = 0.38, *p* = .829). Next, we compared the model with all three-way interaction terms to the same model without the three-way interaction terms. We found that the three-way interaction terms significantly improved model fit (LRT *X*2 (7) = 14.37, *p* = .045). The results are summarized in Table 6.

There was a significant interaction between Group Membership and Response Type such that participants were more likely to punish/reject outgroup members compared to ingroup members and neutral group members, especially in the Punishment Condition compared to the Partner Rejection Condition (Table 1). Indeed, follow-up analyses on the data from the Partner Rejection Condition revealed that the three-way interaction term (Group MembershipXIntentXOutcome) did not improve model fit (LRT *X*2 (2) = 4.04, *p* = 0.13) and, while two-way interaction terms improved model fit (LRT *X*2 (5) = 14.45, *p* = 0.01), we found no interactions with group membership (all OR < 1.11 and > 0.72, *p* > 0.20).

We found a main effect of Group Membership such that participants were more likely to punish/reject outgroup members than ingroup members and neutral group members, but no more likely to punish/reject neutral members compared to ingroup members. In contrast to Study 1, there was a significant interaction between Group Membership and Outcome such that participants were more likely to punish/reject outgroup members compared to ingroup members, especially when there was an equal outcome compared to when there was an unequal outcome.

Additionally, we found a significant three-way interaction between Group Membership, Intent, and Outcome such that the tendency to punish/reject outgroup members more than ingroup members when the outcome was equal compared to when it was unequal was slightly greater when the intent was fair than when it was selfish (Table 7).

Replicating the results of Study 1, we found a significant main effect of Intent such that participants were more likely to punish/reject when Deciders had a selfish intent than when they had a fair intent. Additionally, there was a significant main effect of Outcome such that participants were more likely to punish/reject when there was an unequal outcome than when there was an equal outcome. Once again, there was a significant interaction between Intent and Outcome such that participants were more likely to punish/reject selfish Deciders than fair Deciders, especially when the outcome was unfair than when the outcome was fair (Table 4).

We found a significant main effect of Response Type such that participants rejected partners more often than they punished partners. There was a significant interaction between Response Type and Intent such that participants rejected partners more often than they punished partners, especially when the Decider’s intent was selfish.

# General discussion

Across two studies, we investigated the impact of group membership on two ways of responding to others’ transgressions: punishment and partner rejection. Participants in both studies interacted with other players in an online economic game for real stakes. We assigned participants to groups using a “minimal” groups paradigm (Study 1) and a consequential political position (Study 2). In line with prior work [15], we find that people punish outgroup members more often and more harshly than ingroup members, especially when the partner’s intent was selfish and the outcome was unequal. Extending this work, we find that, in contrast, people are not sensitive to group membership when deciding whether to reject their partner. This effect is robust across both methods of assigning group membership.

Why might punishment be sensitive to group membership in a way that partner rejection is not? One potential explanation centers on the functional differences between punishment and partner rejection. While punishment may serve to help maintain the well-being of the group or define group boundaries, partner rejection may have more immediate and personal consequences for individual well-being. Thus, when deciding to reject a partner or not, people may place more weight on factors that have a more direct impact on their well-being, such as the partner’s intent and the outcome of the interaction, than on group membership, which does not directly impact their well-being. Conversely, punishment may serve to maintain and strengthen group boundaries by imposing more costs on outgroup members than ingroup members. By punishing outgroup members more harshly than ingroup members, people may act in accordance with perceived group norms regarding differential treatment of ingroup versus outgroup members. Ingroup bias in punishment decisions may therefore also serve a group-strengthening function, strengthening participant’s attachment to their ingroup and their detachment from the outgroup. Future work should continue to explore the psychological mechanisms responsible for the difference between punishment and partner rejection in sensitivity to group membership.

Building on prior work on third-party punishment [11], we also found that people were more likely to punish outgroup members compared to neutral individuals. Although we found that people were more likely to punish neutral individuals compared to ingroup members, this difference was not significant. These results suggest that ingroup bias in the context of punishment is driven more by outgroup derogation than ingroup love, consistent with other work on intergroup bias [43][44][45]. However, it is possible that the difference in punishment toward neutral individuals and ingroup members in our study may have been artificially reduced by the fact that in the equal outcome conditions, punishment rates were near floor for all types of group members. Indeed, prior work [11] has found that neutral individuals are punished more than ingroup members. Given this, future work should continue to explore the robustness of differences in punishment between ingroup members and neutral individuals.

***Limitations***

In the present work, we largely explored differences in punishment and partner rejection when they are costless for the responders. Since prior work has demonstrated a consistent influence of group membership on punishment both when afflicting punishment is costly or costless for responders [46][47][48][49], we do not expect that our results for the Punishment Condition would have been different if punishment decisions were costly. However, it is possible that group membership may have a stronger impact on partner rejection when partner rejection decisions are costly. If partner rejection functions primarily to maintain individual well-being, partner rejection decisions may be more sensitive to costs that differentially impact well-being in the context of different social groups. For example, the social costs associated with rejecting a partner from one’s own group may damage one’s reputation with the ingroup, limiting their access to resources or important social networks. The perceived threats of these lost social ties may override the perceived immediate benefits to one’s well-being for rejecting a transgressive partner. On the contrary, rejecting a partner from the outgroup may not pose the same social cost, and thus, would not pose as much of a long-term threat to individual well-being. Thus, the perceived immediate personal benefits of partner rejection may be more persuasive for rejecting an outgroup member. Future work should investigate the influence of group membership on partner rejection decisions when these decisions are costly.

This work is also limited in that we investigated the influence of group membership on only second-party punishment and partner rejection decisions (i.e., those made by the victims of a transgression). The dyadic nature of these interactions may explain why we observed a stronger influence of intent and outcome on punishment and partner rejection than group membership. Indeed, prior work shows that people are especially more likely to punish outgroup members than ingroup members when the punisher is a third-party observer [28][13] and less so when the punisher is the victim [15]. It is unclear how the role of partner rejection decisions might change when made by third-party observers compared to second-party victims. Just as second-party victims seek to maximize their welfare by rejecting partners that have previously harmed them, third-party observers might reject partners that have previously harmed others to maximize their welfare. It is possible that partner rejection decisions made by third-party observers also serve as a form of social signaling, similar to punishment. Given that third-party observers who reject a partner do so based on a harm they observed, rather than a harm they experienced, their rejection may signal group expectations around fair treatment. Future work should clarify the impact of group membership on partner rejection decisions made by third-party observers.

Another limitation of the present work is that we conducted the economic games with samples of online participants, who may be less attentive than participants who participate in economic games in person. A lack of attentiveness could explain why, at an absolute level, intent had a weaker impact on punishment/partner rejection decisions in our studies than in past work using in-person samples of undergraduate participants [34][35]. Follow-up analyses revealed that the impact of intent was stronger among participants who got more comprehension questions correct on their first try, as well as among participants who scored higher on the Cognitive Reflection Test. However, we excluded participants who were unable to answer comprehension checks questions correctly, so this difference in attentiveness should not have impacted participants’ understanding of the task. Additionally, controlling for attentiveness and CRT performance in our analyses did not change the influence of group membership between conditions. Future work should clarify what factors impact the strength of the influence of intent on punishment and partner choice.

**Conclusion**

Group-based psychology pervades our social life. When interacting with another person produces an unequal outcome for oneself, or suggests that the other person was behaving selfishly, decisions to punish or reject the person are differentially influenced by the person’s group membership status. Across minimal and consequential manipulations of group membership, we find that people are more likely to punish outgroup members than ingroup members, but no more likely to reject them. These results shed light on the underlying functions of punishment and partner rejection and expand our understanding of the role that groups play in our social decision-making.

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**Author Contributions**

Trystan Loustau was responsible for writing the main manuscript. Jacob Glassman conducted the main data analyses and sensitivity analyses. Justin W. Martin designed the original study, collected the study data, drafted the original version of the manuscript, and conducted the main and supplementary analyses. Liane Young provided feedback on the study design and manuscript. Katherine McAuliffe-provided feedback on the study design and manuscript.

**Data availability**

Data and analysis scripts for all studies can be found at <https://osf.io/cjs8q/>. For both studies, we report all measures, manipulations and exclusions. The pre-registration for Study 1 can be found at <http://aspredicted.org/blind.php?x=j92vh7>. The pre-registration for Study 2 can be found at <http://aspredicted.org/blind.php?x=56xa9a>. All materials can be found in the Supplementary Materials. All procedures for both studies were approved by the Boston College Institutional Review Board.

**Competing Interests Statement**

We have no competing interests to disclose.

**Figure Legends**

**Figure 1. Legend**

1. *Mean responses in the Punishment and Partner Rejection Conditions in Study 1*. Plotted are participants’ mean responses across the Intention, Outcome, and Group membership conditions for the Punishment (Panel A) and Partner Choice (Panel B) conditions. Error bars are SEM. Note that possible punishment values ranged from -$0.30 to $0.30; we plot a restricted range here for visualization purposes.
2. *Mean responses in the Punishment and Partner Rejection conditions in Study 2.* Plotted are participants’ mean responses across the Intention, Outcome, and Group membership conditions for the Punishment (Panel A) and Partner Choice (Panel B) conditions. Error bars are SEM.

**Tables**

**Table 1**

*Four Conditions (Studies 1-2)*

|  |  |  |  |
| --- | --- | --- | --- |
|  | | **Decider’s Intent** | |
| Option A (selfish) | Option B (fair) |
| **Outcome** | Equal | Selfish Intent | Equal Outcome | Fair Intent | Equal Outcome |
| Unequal | Selfish Intent | Unequal Outcome | Fair Intent | Unequal Outcome |

**Table 2**

*Linear Mixed Effects Regression on Punishment (Study 1)*

|  |  |  |  |  |
| --- | --- | --- | --- | --- |
| **Effect** | ***B*(SE)** | **t** | **p** | **95% CI** |
| Intent | -0.41(0.05) | -8.58 | <.001 | (-0.50, -0.31) |
| Outcome | -0.73(0.05) | -15.98 | <.001 | (-0.82, -0.64) |
| Group Membership | 0.06(0.02) | 3.81 | <.001 | (0.03, 0.09) |
| IntentXOutcome | -0.09(0.03) | -3.48 | <.001 | (-0.14, -0.04) |
| IntentXGroup Membership | -0.02(0.03) | -0.65 | .515 | (-0.07, 0.03) |
| OutcomeXGroup Membership | -0.004(0.03) | -0.14 | .888 | (-0.05, 0.05) |

**Table 3**

*Mean Engagement in Punishment or Partner Rejection (Studies 1-2)*

|  |  |  |  |  |
| --- | --- | --- | --- | --- |
| **Study** | **Response Type** | **Mean or Proportion (SEM)** | | |
| Ingroup | Neutral | Outgroup |
| 1 | Punishment (Mean) | -0.01 (0.01) | NA | -0.03 (0.01) |
| Partner Rejection (Proportion) | 0.53 (0.01) | NA | 0.52 (0.01) |
| 2 | Punishment (Mean) | 0.29 (0.01) | 0.33 (0.01) | 0.42 (0.01) |
| Partner Rejection (Mean) | 0.44 (0.01) | 0.45 (0.01) | 0.47 (0.01) |

**Table 4**

*Impact of Intent and Outcome on Punishment and Partner Rejection (Studies 1-2)*

|  |  |  |  |  |
| --- | --- | --- | --- | --- |
| **Study** | **Intent** | **Outcome** | **Response Type** | |
| Punishment | Partner Rejection |
| 1 | Fair | Equal | 0.10 (0.01) | 0.97 (0.01) |
| Fair | Unequal | -0.05 (0.01) | 0.50 (0.02) |
| Selfish | Equal | 0.02 (0.01) | 0.54 (0.02) |
| Selfish | Unequal | -0.14 (0.01) | 0.07 (0.01) |
| 2 | Fair | Equal | 0.10 (0.01) | 0.04 (0.01) |
| Fair | Unequal | 0.48 (0.02) | 0.45 (0.02) |
| Selfish | Equal | 0.18 (0.01) | 0.44 (0.02) |
| Selfish | Unequal | 0.63 (0.02) | 0.87 (0.01) |

*Note.* Punishment scores represent Means and SEMs. Partner Rejection scores represent Proportions and SEMs in Study 1 and Means and SEMs in Study 2.

**Table 5**

*Logistic Regression on Partner Rejection (Study 1)*

|  |  |  |  |  |
| --- | --- | --- | --- | --- |
| **Effect** | **OR(SE)** | **z** | **p** | **95% CI** |
| Intercept  Intentions | 1.31 (0.07)  0.06 (0.15) | 3.737  -19.89 | <.001  <.001 | (1.14, 1.52)  (0.04, 0.07) |
| Outcomes | 0.05 (0.15) | 20.901 | <.001 | (0.04, 0.06) |
| Group Membership | 1.09 (0.15) | 0.562 | .574 | (0.8, 1.45) |
| IntentionsXOutcomes  Group Membership XIntent  Group Membership X Outcome  IntentionsXOutcomeXGroup Membership | 2.05 (0.29)  0.83 (0.29)  0.85 (0.29)  1.20 (0.58) | 2.472  -0.629  -0.545  0.309 | .013  .530  .586  .758 | (1.18, 3.69)  (0.47, 1.47)  (0.48, 1.51)  (0.38, 3.80) |

**Table 6**

*Mixed Effects Logistic Regression Analysis (Study 2)*

|  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- |
| **Effect** | **B (SE)** | **OR** | **SE** | **z** | **p** | **95% CI** |
| Intercept | -1.23 (0.10)\*\*\* | 0.29 | 0.10 | -12.33 | <.001 | (0.24, 0.35) |
| Group Membership: Neutral | 0.26 (0.10)\* | 1.29 | 0.10 | 2.54 | .011 | (1.06, 1.58) |
| Group membership: Outgroup | 0.78 (0.10)\*\*\* | 2.18 | 0.10 | 8.10 | < .001 | (1.84, 2.64) |
| Intent | 2.09 (0.16)\*\*\* | 8.11 | 0.16 | 13.45 | < .001 | (6.23, 11.02) |
| Outcome | 3.31 (0.16)\*\*\* | 27.49 | 0.16 | 20.94 | < .001 | (20.49, 39.25) |
| Response Type | -1.08 (0.19)\*\*\* | 0.34 | 0.19 | -5.76 | < .001 | (0.24, 0.49) |
| Intent X Group member: Netural | -0.05 (0.21) | 0.95 | 0.21 | -0.25 | .801 | (0.65 1.36) |
| Intent X Group member: Outgroup | -0.22 (0.20) | 0.80 | 0.20 | -1.12 | .262 | (0.55 1.14) |
| Outcome X Group member: Netural | -0.08 (0.21) | 0.92 | 0.21 | -0.39 | .693 | (0.63, 1.39) |
| Outcome X Group member: Outgroup | -0.66 (0.19)\*\*\* | 0.52 | 0.19 | -3.39 | < .001 | (0.35, 0.75) |
| Intent X Outcome | -0.85 (0.31)\*\* | 0.43 | 0.31 | -2.76 | .006 | (0.24, 0.75) |
| Response type X Group member: Netural | 0.22 (0.18) | 1.25 | 0.18 | 1.22 | .221 | (0.89, 1.79) |
| Response type X Group member: Outgroup | 0.75 (0.18)\*\*\* | 2.13 | 0.18 | 4.30 | < .001 | (1.48, 3.03) |
| Response type X Intent | -1.94 (0.28)\*\*\* | 0.14 | 0.28 | -6.98 | < .001 | (0.08, 0.24) |
| Response type X Outcome | 0.13 (0.29) | 1.14 | 0.29 | 0.45 | .652 | (0.64, 2.05) |
| Intent X Outcome X Group member: Netural | 0.37 (0.42) | 1.45 | 0.42 | 0.89 | .372 | (0.64, 3.19) |
| Intent X Outcome X Group member: Outgroup | 0.98 (0.39)\* | 2.66 | 0.39 | 2.52 | .012 | (1.27, 5.58) |
| Intent X Response type X Group member: Netural | -0.20 (0.37) | 0.82 | 0.37 | -0.54 | .590 | (0.40, 1.62) |
| Intent X Response type X Group member: Outgroup | 0.02 (0.36) | 1.02 | 0.36 | 0.05 | .959 | (0.50 2.05) |
| Outcome X Response type X Group member: Neutral | -0.12 (0.39) | 0.88 | 0.39 | -0.32 | .750 | (0.43, 1.92) |
| Outcome X Response type X Group member: Outgroup | -0.43 (0.37) | 0.65 | 0.37 | -1.15 | .250 | (0.31 1.36) |
| Intent X Response type X Outcome | 0.60 (0.32) | 1.83 | 0.32 | 1.89 | .058 | (0.99, 3.89) |

*Notes: \*\*\*p < 0.001, \*\*p < 0.01, \*p < 0.05*

**Table 7**

*Mean Engagement in Punishment or Partner Rejection (Study 2)*

|  |  |  |  |  |
| --- | --- | --- | --- | --- |
|  |  | **Mean (SEM)** | | |
| **Intent** | **Outcome** | Ingroup | Neutral | Outgroup |
| Fair | Equal | 0.04 (0.01) | 0.05 (0.01) | 0.12 (0.01) |
| Fair | Unequal | 0.44 (0.02) | 0.47 (0.02) | 0.49 (0.02) |
| Selfish | Equal | 0.27 (0.02) | 0.29 (0.02) | 0.37 (0.02) |
| Selfish | Unequal | 0.71 (0.02) | 0.75 (0.02) | 0.79 (0.02) |

**Table 8**

*Mean Group Identification Ratings (Studies 1 and 2)*

|  |  |  |  |
| --- | --- | --- | --- |
|  | **Mean (SEM)** | | |
| **Study** | Ingroup | Neutral | Outgroup |
| Study 1 | 62.48 (0.91) | NA | 40.97 (0.71) |
| Study 2 | 72.44 (0.87) | 40.17 (0.87) | 29.58 (0.96) |
|  |  |  |  |

**Supplementary**

**Analysis Approach for Study 1**

Including random slopes generated significant convergence issues. To address convergence issues, we ran standard regression models run with Firth bias correction and found that all significant fixed effects were supported or a greater number of significant effects, indicating that the random intercept models did not inflate type 1 error rate.

**Order Effects in Study 1**

We found a complicated pattern based upon the order in which participants encountered ingroup and outgroup Deciders. Participants weighted outcomes more heavily for the first Decider they encountered, staying if this decider caused a fair outcome and switching if they caused an unfair outcome. Importantly, we continued to find no overall influence of group membership, nor any interaction between group membership and intentions or outcomes.

**Exploratory Measures in Study 2**

**Playing as Decider.** Participants then played the game three more times, now playing as the Decider. Participants were provided with the same set of instructions but now described from the Decider’s perspective. Participants again answered a series of 7 comprehension questions, presented only once and not used in screening data quality. Participants played the game once with each type of player, with order counterbalanced across participants.

**CRT, attention check questions and debrief.** Participants next completed the 7-item CRT (Pennycook & Rand, 2020). Then they completed six questions designed to assess the degree to which the other player’s unfair choice and the unfair outcome were perceived as negative. For each type of player, participants were asked “When the partner who identified as [type] chose the unfair option (Option A), how negative did this make you feel?” and “When the partner who identified as [type] received the entire $1, how negative did this make you feel?”, with both questions asked on 7-point scales anchored at 1 = “Not at all negative” and 7 = “Extremely negative”. Finally, participants completed the same set of attention check questions as in Study 1 and were debriefed.

**Exploratory Analyses in Study 2**

*Firth Bias Correction*

Inspection of the main mixed-effects logistical regression model in Study 2 and the distribution of our response variable indicated that some cells of our design had very few responses. For instance, when paired with an Ingroup Decider, in the case where the Decider had a fair intent and produced a fair outcome, only 8/300 responses (0.3%) were to reject the Decider, and 13/300 responses (0.4%) were to engage in punishment. Similarly, when paired with a Neutral Decider, in the same case, only 9/300 responses (0.3%) were to reject the Decider and 22/300 responses (0.7%) were to engage in punishment. Typical procedures to address convergence issues were not able to overcome this issue of rare events. Because mixed-effects models only including random intercepts can yield inflated type 1 error rates (Barr, Levy, Scheepers, & Tily, 2013; but see Douglas Bates, Kliegl, Vasishth, & Baayen, 2015; Matuschek, Kliegl, Vasishth, Baayen, & Bates, 2017), we additionally model our results using standard logistic regression (without random effects) while performing Firth bias correction (Firth, 1993), which reduces small-sample bias and yields definite estimates. Firth bias correction was implemented using the brglm2 package (Kosmidis, 2020). In all cases, these models either found the same pattern of statistical results or yielded a larger number of significant fixed effects, suggesting that our random effects models were not inflating type 1 error rates.

*Classification Analyses*

Additionally, in order to facilitate comparison to the results of Study 1, we classified each participant as being sensitive to intentions, outcomes and group membership (non-mutually exclusive) and then compared these percentages across the Punishment and Partner Rejection conditions (see Supplemental Material). We used the same approach as in Study 1, classifying participants as sensitive to a feature (intentions, outcomes and group membership) if, holding the other two features constant, their response differed based upon the levels of that feature (ignoring the direction of that difference).

As in Study 1, we explored the influence of a number of additional variables on this overall pattern of results, especially the interaction between group membership and response type (see Supplemental Material). First, we found that these results were unchanged if insensitive participants were excluded, with the exception that we found additional interactions between group membership, intentions and outcomes, and between response type, intentions and outcomes. Similarly, we continued to find an interaction between response type and group membership when including ingroup identification ratings in our models. However, we now additionally found 2-way interactions between ingroup identification and group membership, outcomes and intentions, as well as a marginal interaction between ingroup identification and response type. Including a predictor for which Decider participants interacted with first yielded no influence of Decider order on the influence of group membership. We continued to find an interaction between response type and group membership, though we found additional interactions between Decider order and intentions and Decider order and outcomes. Including a predictor for attentiveness (as measured by first-try performance on comprehension questions), we found an interaction between attentiveness, outcomes and group membership; attentiveness, response type and outcomes; and between attentiveness, response type and intentions. However, attentiveness did not influence the interaction between response type and group membership, which remained significant even when the role of attentiveness was accounted for. Finally, including CRT performance in our models yielded a marginal interaction between CRT performance, response type and group membership. Inspecting means, this interaction reflected a greater tendency to differentiate outgroup partners from ingroup and neutral partners amongst those higher on CRT performance, relative to those lower on CRT performance. In addition, those higher on CRT performance were less likely to engage in punishment overall and more likely to reject the Decider, relative to those lower on CRT performance. We note, however, that strong conclusions should not be drawn from these results, given their marginal nature. We also found interactions between CRT performance, response type and outcomes and between CRT performance, response type and intentions. However, we continued to find an interaction between response type and group membership when the influence of CRT performance was accounted for. Thus, while these additional factors influenced the role that other factors play in some cases, the interaction between response type and group membership was consistently significant and only ever modulated in the case of the marginal interaction with CRT performance. In total, we consistently found support for the idea that group membership has a strong influence on punishment and has almost no influence on partner choice decisions.

*Participants as the Decider*

When analyzing participants’ decisions when playing as the Decider, responses are binary (Option A versus Option B) and so we model these data using mixed-effects logistic regression, including a random intercept for participants and again taking a model comparison approach.

Finally, as in Study 1, we investigated how participants’ decisions as a Decider were influenced by group membership and whether they might be subject to punishment or partner choice (see Supplemental Material). To summarize, we found that participants were most likely to choose the fair option when the Responder was an ingroup member, slightly less likely to choose the fair option when paired with a neutral individual, and least likely to choose the fair option when paired with an outgroup member, in both response conditions. In Study 2 we no longer found that response type influenced these decisions. We again found that the influence of Responder group membership was stronger when individual differences in ingroup identification were taken into account, such that those highest on ingroup bias were more likely to favor ingroup Responders over outgroup or neutral Responders.

Judgments

We also found no difference in judgment of ingroup and neutral partners (OR = 1.21, SE = 0.11, *z =* 0.69, *p* = 0.49, 95% CI = 0.87–1.34) or neutral and outgroup partners (OR = 1.17, SE = 0.11, *z =* 1.44, *p* = 0.15, 95% CI = 0.95–1.45), though we did find a weak difference in judgment of ingroup and outgroup partners (OR = 1.26, SE = 0.11, *z =* 2.14, *p* = 0.03, 95% CI = 1.02–1.57).